

# Working Horizon and Labour Supply: the Effect of Raising Minimum Retirement Age on Middle-aged Individuals

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- Aging is major concern for OECD countries. Pension reforms increasing legal retirement age widely adopted
- Literature mostly focused on the LS response by older individuals close to eligibility (direct effect)
- **This paper:** effects on individuals far from the eligibility, whose working horizon increases→ e.g., retiring in 7 rather than 2 years  
**Any forward looking effect on their LS?**

# Why do we expect an effect?

- Non-myopic individuals, possible channels:
  - ① Pension wealth  $\downarrow \rightarrow$  LS  $\uparrow$
  - ② Need to work longer but expecting worse conditions when old  $\rightarrow$  LS  $\uparrow$
  - ③ Working period  $\uparrow \rightarrow$  returns to work  $\uparrow$ , search and LS  $\uparrow$
- Participation rates particularly low for women  $\rightarrow$  Unutilized capacity

# Scant literature on forward looking behaviour

- Hairault et al. (2010): French reform  $\uparrow$  FRA by 1 year; men 57+
- Geyer and Welteke (2019): German reform  $\uparrow$  ERA by 3 years; 58-59 y.o. women
- We look at a bigger reform implemented in Italy
  - ① sizeable  $\uparrow$  in FRA up to 7 years
  - ② unexpected and well-understood
  - ③ FRA  $\uparrow$  heterogeneously depending on observable dimensions

# What do we do?

Exploiting the heterogeneity in the increase in FRA we estimate (DD framework):

- **Forward looking effect:** LS responses of middle aged individuals
- **Spillover effects on the their partner's LS**
- Well-known **direct effect** on close-to-eligibility individuals and their partners

# Preview of the main findings

- Forward looking effect:
  - **Positive on LS for middle-aged women** not for men
  - Stronger for lower educated women
  - Positive spillovers into husbands' LS
- Direct effect: positive on LS, comparable with other studies

# The Treatment and the Data

- Two types of pension schemes in Italy for full retirement:
  - ① Old age: minimum age required
  - ② Seniority: given # of years of paid contribution to the Social Security system



- Stricter eligibility requirements for both schemes
  - ① Old age: RA ↑ from 60 to 67 y.o. for women, from 65 to 67 for men
  - ② Seniority:
    - Before: i) quota (60 y.o. + 35 years of contribution) or ii) 40 years of contribution
    - After: i) quota abolished; ii) 42 (women), 43 (men) years of contribution

# Distance to retirement and its change

- $MRA_t = \min(OLD\ AGE_t, SENIORITY_t)$ ; rules at  $t$ . Hp: from  $t$  onward assume individuals work every period

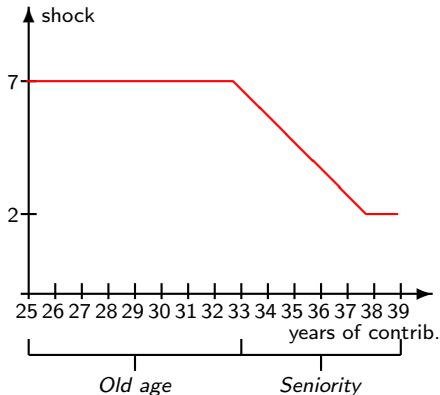
	2014	
	Seniority	Old
Women, 58 years old		
Maria, $C = 38$	<b>62</b>	67
Valeria, $C = 26$	73	<b>67</b>

- MRA changes depending on: **age, accrued years of contribution, gender**; each combination is a given **cell**  $q$
- **Time invariant** measure of policy shock defined at cell  $q$  level:

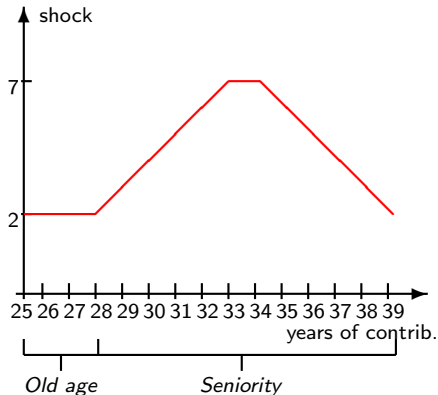
$$T_q = MRA_q^{post} - MRA_q^{pre}$$

# Heterogeneous change in MRA, years of contribution

Woman 58 y.o.

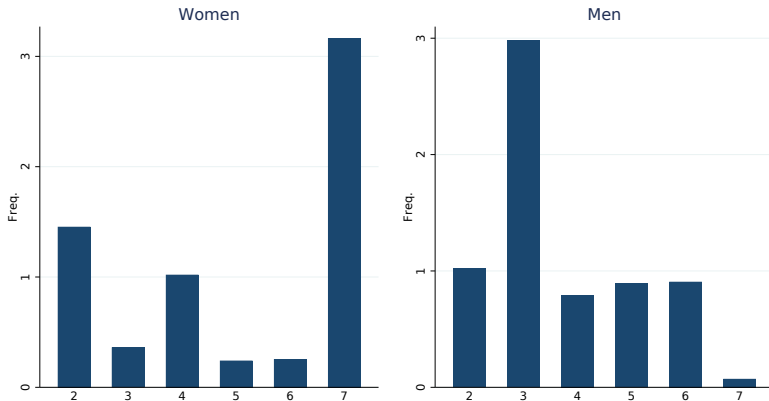


Man 58 y.o.



- The Italian Survey of Household Income and Wealth (SHIW) from 2004 to 2016
  - biannual survey on 8,000 households per wave
  - information on accrued years of contribution
  - information on partners (family level analysis)
  - information on all different source of income
  - information on expected retirement age
- More results with administrative data from Social Security accounts

# Distribution of the shock $T_q$ , by gender



**Source:** SHIW, 2012-2016.

**Note:** women (men) aged between 45 and 59 (45-64), with at least 10 (20) accrued years of contribution, eligible to retire neither before nor after the reform. Data are at the individual level.

- Women have less years of paid contribution than men
- Different change in rules for old age scheme

# The Forward Looking Effect: individual level analysis

# Main regression equation, individual level analysis

**Separately for men and women by age groups in 2004-2016:**

$$Y_{iqt} = \beta_1 T_q * post2011_t + \beta_2 X_{iqt} + \alpha_t + \alpha_q + \epsilon_{iqt}$$

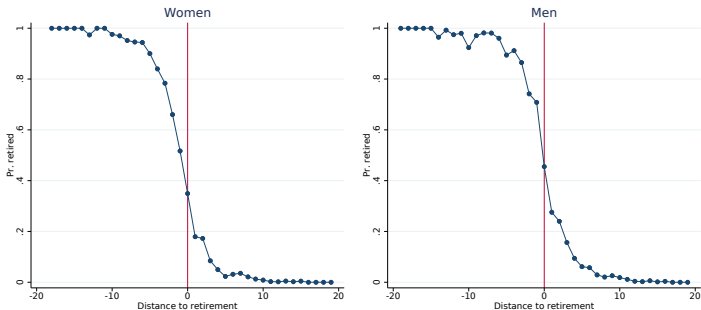
- $Y_{iqt}$  dummy = 1 if the individual  $i$  in cell  $q$  is active at  $t$
- $T_q$  time invariant measure of policy shock: the cross-sectional variation in MRA occurs at level  $q$  (age, # years of contribution)
- $\alpha_q$  cell  $q$ -fixed effects;  $\alpha_t$  year-fixed effects
- $X_{iqt}$  are controls (marital status, region, sector)
- $\epsilon_{iqt}$  error term
- standard errors clustered at the  $q$  level

for individuals **not eligible to retire even before the reform but reasonably close to retirement** [▶ Descriptive statistics](#)

# Three identifying assumptions

## 1. MRA defines the length of the working horizon

Prob. of retiring and distance to MRA ( $DMRA = MRA_t - age_t$ )



Source: SHIW, 2008-2016.

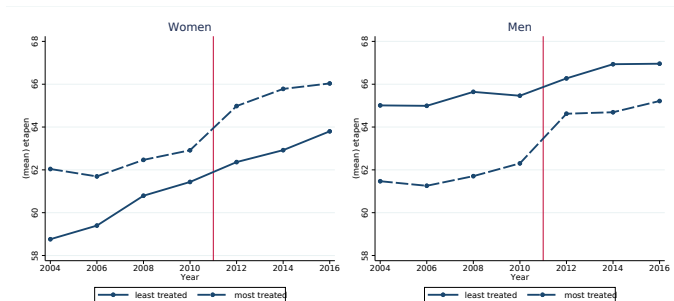


# Three identifying assumptions

## 2. Expected retirement age changes according to the rules in place

Expected RA increases more for treated than for control individuals

Expected RA by exposure to the shock



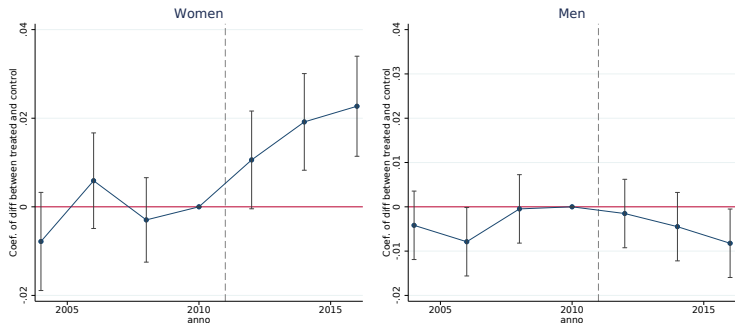
**Source:** SHIW, 2008-2016. **Note:** most treated are those whose shock to MRA is larger than 4 years.

# Three identifying assumptions

## 3. Parallel trends in the outcome

Evolution of the difference in the probability of being active between more and less exposed individuals

$$Y_{igt} = \sum_{r=2004}^{2016} \gamma_r (T_q * \delta_r) + \gamma_1 X_{igt} + \delta_q + \delta_t + \eta_{igt}$$



**Note:** The graphs plots the coefficients  $\gamma_r$  and the corresponding 95% confidence intervals.

# Positive effects on female participation at all ages

	Women			Men		
	55-59	50-54	45-49	55-64	50-54	45-49
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Participation</i>						
T*post2011	0.033*** (0.006)	0.015*** (0.004)	0.011*** (0.004)	-0.002 (0.004)	-0.003 (0.002)	0.005* (0.003)
<i>Unemployment</i>						
T*post2011	0.012** (0.006)	0.009*** (0.004)	0.006** (0.003)	-0.004 (0.006)	-0.001 (0.005)	0.009 (0.008)
<i>Employment</i>						
T*post2011	0.022*** (0.006)	0.006 (0.004)	0.004 (0.005)	0.001 (0.007)	-0.002 (0.006)	-0.004 (0.009)
N	2456	3332	3091	3577	3856	3194

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

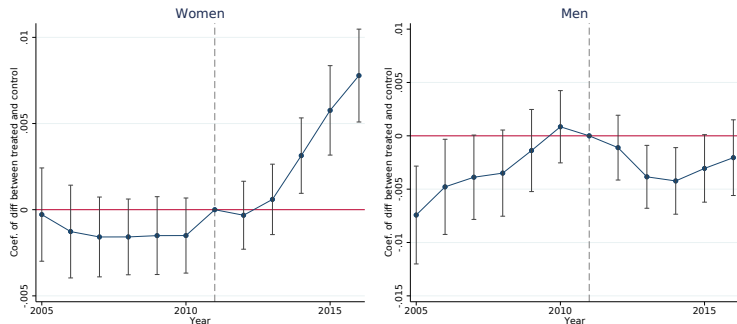
- More full-time less part-time ► type
- Income and savings  $\uparrow$  for women 55-59 y.o. ► income

- Stronger effects for low educated individuals ▶ education
- Unemployment effects also in high demand regions ▶ high vacancy

More results with the use of  
administrative panel data

- Social Security Records: 6.6% random sample of employees in the private sector
  - ① Employer-employees data (wages, contract, occupation), 2005 to 2016
  - ② Accrued years of contribution recorded since the first year of work
  - ③ Information on retirees (year of retirement)
- Analysis run as before
  - Pros:
    - ① Recorded number of years of contribution - no self-reported as in SHIW
    - ② Drivers of employment: more newly employed or less separated?
  - Cons:
    - ① No family dimension
    - ② Only employment

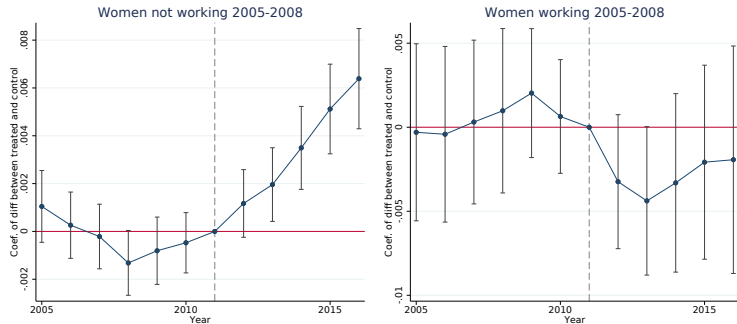
# Similar employment effects



**Source:** Social security administrative records.

**Note:** Pre-reform years: 2005-2011, post-reform years: 2012-2016. The omitted year is 2011. Sample: individuals with at least 10 (for women) or 20 (for men) and less than 40 accrued years of contribution; women aged 50-59; men aged 55-64, not eligible to retire either before or after the 2012 pension reform.

# More newly employed or less separated?



**Source:** Social security administrative records.

**Note:** Pre-reform years: 2005-2011, post-reform years: 2012-2016. The omitted year is 2011. Sample: individuals with at least 10 (for women) or 20 (for men) and less than 40 accrued years of contribution; women aged 50-59; men aged 55-64, not eligible to retire either before or after the 2012 pension reform.

► More results



Spillover effects on partners

# Main regression equation, family level analysis

We estimate the following equation for husbands and wives in 2004-2014:

$$Y_{jq_s q_{s'} t}^s = \beta_1^s T_{q_{s'}}^{s'} * post2011_t + \beta_2^s T_{q_s}^s * post2011_t + \beta_2^s W_{jt} + \alpha_{q_{s'}}^s + \alpha_{q_s}^s + \alpha_t^s + u_{jq_s q_{s'} t}^s$$

- $Y_{jq_s q_{s'} t}^s$  dummy = 1 if spouse  $s$  in household  $j$  belonging to cell  $q_s$ , whose partner  $s'$  belongs to the age-contribution cell  $q_{s'}$
- $T_{q_{s'}}^{s'}$  time invariant measure of the policy shock for partner  $s'$
- $T_{q_s}^s$  time invariant measure of the policy shock of individual  $s$
- $\alpha_{q_{s'}}^s$  and  $\alpha_{q_s}^s$  fixed effects for cell  $q_{s'}$  and  $q_s$ ;  $\alpha_t^s$  year-fixed effects
- $W_{jt}$  controls at individual and household level
- $u_{jq_s q_{s'} t}^s$  is an error term

**Partner  $s'$  not eligible to retire under pre-reform rules; partner  $s$  no restrictions**

► Descriptive statistics

# Evolution of the difference in the spouse's probability of being active between more and less exposed individuals



**Source:** SHIW, from 2004 to 2016.

**Note** The graphs test the parallel trend assumption by plotting the coefficients  $\zeta_r^s$  and the corresponding 95% confidence interval obtained from estimating equation  $Y_{iq_{s,t}}^s = \sum_{r=2004}^{2014} \zeta_r^s (T_{q_{s,t}} * \alpha_r^s) + \beta^s W_{iq_{s,t}} + \alpha_{q_{s,t}}^s + \alpha_t^s + \alpha_{q_{s,t}}^s + u_{iq_{s,t}}^s$ . Pre reform years: 2004-2006-2008-2010, post-reform years: 2012-2014-2016.

# Cross elasticities among partners

	Shock to wife MRA		Shock to husband MRA	
	on wife (1)	on husband (2)	on husband (3)	on wife (4)
<i>Participation</i>				
T wife*post2011	0.022*** (0.003)	0.018** (0.007)		
T husb*post2011			-0.002 (0.003)	-0.011 (0.010)
<i>Unemployment</i>				
T wife*post2011	0.013*** (0.003)	0.006 (0.004)		
T husb*post2011			-0.001 (0.005)	-0.008 (0.006)
<i>Employment</i>				
T wife*post2011	0.010** (0.004)	0.012 (0.008)		
T husb*post2011			-0.001 (0.006)	-0.003 (0.011)
N	5326	5326	3819	3819

**Notes:** Additional controls: year and cell  $q_s$ , and  $q_s$  fixed effects, region and sector fixed effects, age difference across partners (also squared) and difference in distance to retirement across partners, partner  $s$  change in distance to retirement. The sample in columns 1 and 2 (3 and 4) consists of couples where the wives (husbands) are not eligible for a public pension either before and after the reform and have accrued more than 9 (19) and less than 40 years of contribution. Robust standard errors in parentheses.

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

[► full specification](#)

- Husbands respond to their wives' shock
  - ① **The effect comes from men eligible to retire but postpone retirement**
  - ② Stronger effects for low educated ▶ education
- Wives do not respond since men are not reacting to begin with

# The Direct Effect

# The direct effect on individuals eligible before the reform

- Look at individuals close to the eligibility threshold before the reform:  
→ RDD framework
- Diff-in-disc: comparing the probability of being active in the pre and in the post-reform period of individuals around the 2010 eligibility threshold ► Specific.
- Direct effect: in post-reform years larger proportion of active individuals around the 2010 eligibility threshold than in pre-reform years
- Positive effects both for women and for men, no cross-partner effects  
► Individuals ► Partners

# Summing up on the magnitude

- Coefficient of the direct effect is about 8 times larger than that of forward-looking effect (0.24 against 0.03 for women)
- But:
  - Within household spillovers stronger for forward looking (FL) effect than for direct effect (labour supply of older men less elastic, most of them already retired)
  - Share of individuals affected by the FL much larger
- FL effect contributed by one third to the increase in Italian female LFPR after 2012



- Effects of delaying MRA on labour supply go beyond those individuals eligible to retire under the previous rules
- Positive effects on women participation at different age, not on men
- The effect on wives has spillovers on their husbands' participation
- Effect stronger for women less attached to the labour market
- Forward looking effects are quantitatively large; need to be considered when evaluating a pension reform

# Thanks for the attention!

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# Seniority/early pension eligibility

Year	Private & Public		Self-employed	
	A, C, Q	only C	A, C, Q	only C
<i>Before Fornero reform</i>				
2007	57, 35	39	58, 35	40
2008	58, 35	40	59, 35	40
2009-2010	59, 35, 95	40	60, 35, 96	40
2011	60, 35, 96	40	61, 35, 97	40
2011-2012	60, 35, 96	40	61, 35, 97	40
2013 onwards	61, 35, 97	40	62, 35, 98	40
<i>After Fornero reform</i>				
2012- (men)		43		43
2012- (women)		42		42

**Notes:** A stands for age, C for number of years of contribution,  $Q = A + C$  is the so-called "quota", the sum of age and years of contribution must be larger or equal than Q to have retirement eligibility. Independently from actual age, retirement eligibility is also granted when the number of years of contribution is sufficiently high (39 in 2007, 40 in the following years).

# Descriptive statistics

	Individual level analysis							
	Women				Men			
	All 45-59 [1]	Sample not elig [2]	Control $T_q < 7$ [3]	Treated $T_q \geq 7$ [4]	All 45-64 [5]	Sample not elig [6]	Control $T_q < 4$ [7]	Treated $T_q \geq 4$ [8]
Age	51.665 (4.280)	51.335 (4.108)	50.802 (3.815)	51.721 (4.267)	53.877 (5.728)	52.107 (4.626)	52.172 (4.925)	52.019 (4.179)
Y. contrib.	15.853 (13.343)	23.950 (7.696)	30.476 (4.257)	19.227 (5.986)	28.120 (11.331)	28.707 (5.195)	29.502 (5.401)	27.617 (4.685)
Married	0.762	0.718	0.729	0.710	0.830	0.840	0.836	0.845
High edu	0.497	0.616	0.663	0.582	0.484	0.538	0.467	0.636
If children	0.660	0.632	0.605	0.651	0.640	0.617	0.626	0.605
Active	0.583	0.881	0.976	0.812	0.795	0.979	0.973	0.989
Unempl	0.043	0.044	0.022	0.060	0.071	0.050	0.052	0.048
Part time	0.102	0.147	0.144	0.150	0.022	0.017	0.017	0.018
Perm. contr	0.401	0.641	0.806	0.521	0.500	0.674	0.655	0.700
Log(wage net)	9.522 (0.568)	9.560 (0.548)	9.669 (0.425)	9.450 (0.629)	9.799 (0.480)	9.848 (0.428)	9.838 (0.413)	9.860 (0.446)
Observations	16157	9037	3853	5184	19313	10732	6178	4554

**Notes:** Men and women with at least 35 year old and 8 years of accrued contributions. Women (men) are treated if experienced a shock to distance to minimum retirement of  $> 7$  ( $\geq 4$ ) years after 2012 reform. Years 2004 and 2010. Continuity w.l.=y. contrib/age. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

# Results: type of employment

	Women			Men		
	55-59	50-54	45-49	55-64	50-54	45-49
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Full-time employment</i>						
T*post2011	0.024*** (0.008)	0.015** (0.007)	0.018** (0.008)	-0.000 (0.007)	0.004 (0.007)	-0.010 (0.012)
<i>Part-time employment</i>						
T*post2011	-0.003 (0.007)	-0.009 (0.006)	-0.016** (0.007)	-0.001 (0.003)	0.000 (0.004)	0.004 (0.006)
<i>Permanent employment</i>						
T*post2011	0.014* (0.007)	0.002 (0.006)	0.002 (0.006)	0.004 (0.007)	-0.010 (0.008)	-0.020 (0.014)
N	2456	3332	3091	3577	3856	3194

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

# Results: by benefit calculation

	Women			Men		
	55-59	50-54	45-49	55-64	50-54	45-49
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Participation</i>						
T*post2011	0.020*	0.019**	0.017**	-0.006	-0.005	-0.001
	(0.011)	(0.009)	(0.008)	(0.009)	(0.004)	(0.008)
T*post2011*DB	0.053	-	-	0.001	0.001	-
	(0.040)	-	-	(0.008)	(0.013)	-
<i>Unemployment</i>						
T*post2011	0.013	0.018**	0.010	0.001	0.003	0.006
	(0.014)	(0.008)	(0.006)	(0.016)	(0.015)	(0.019)
T*post2011*DB	-0.003	-	-	-0.011	-0.026	-
	(0.028)	-	-	(0.015)	(0.027)	-
<i>Employment</i>						
T*post2011	0.006	0.002	0.007	-0.007	-0.008	-0.006
	(0.015)	(0.010)	(0.011)	(0.018)	(0.014)	(0.023)
T*post2011*DB	0.056*	-	-	0.012	0.027	-
	(0.032)	-	-	(0.016)	(0.029)	-
N	2453	3332	3090	3576	3856	3193

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. High edu is a dummy equal to 1 if individuals obtained at least the secondary school degree. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ ,

# Results: by education level

	Women				Men	
	55-59 (1)	50-54 (2)	45-49 (3)	55-64 (4)	50-54 (5)	45-49 (6)
<i>Participation</i>						
T*post2011	0.053*** (0.009)	0.027*** (0.007)	0.017** (0.008)	-0.005 (0.007)	-0.005 (0.003)	0.005 (0.007)
T*post2011*high edu	-0.041*** (0.010)	-0.023*** (0.009)	-0.009 (0.009)	0.009 (0.009)	0.005 (0.005)	-0.001 (0.008)
<i>Unemployment</i>						
T*post2011	0.020** (0.010)	0.021*** (0.007)	0.010 (0.006)	-0.006 (0.011)	-0.001 (0.010)	0.008 (0.018)
T*post2011*high edu	-0.011 (0.012)	-0.018** (0.007)	-0.005 (0.007)	0.005 (0.012)	0.000 (0.011)	0.001 (0.020)
<i>Employment</i>						
T*post2011	0.033*** (0.012)	0.006 (0.009)	0.007 (0.010)	0.001 (0.012)	-0.004 (0.009)	-0.004 (0.020)
T*post2011*high edu	-0.029** (0.013)	-0.005 (0.011)	-0.004 (0.012)	0.004 (0.015)	0.005 (0.011)	-0.001 (0.022)
N	2453	3332	3090	3576	3856	3193

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. High edu is a dummy equal to 1 if individuals obtained at least the secondary school degree. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ ,

\*  $p < 0.1$ .



# Results: demand or supply?

	Women		Men			
	55-59 (1)	50-54 (2)	45-49 (3)	55-64 (4)	50-54 (5)	45-49 (6)
<i>Participation</i>						
T*post2011	0.037*** (0.010)	0.010 (0.008)	0.011 (0.008)	0.001 (0.006)	-0.000 (0.004)	0.005 (0.005)
T*post2011*high lambda	-0.004 (0.013)	0.011 (0.011)	-0.000 (0.009)	-0.009 (0.009)	-0.006 (0.005)	0.000 (0.007)
<i>Unemployment</i>						
T*post2011	0.006 (0.009)	0.011 (0.008)	0.014* (0.008)	-0.006 (0.009)	-0.007 (0.008)	0.000 (0.014)
T*post2011*high lambda	0.002 (0.012)	-0.002 (0.009)	-0.011 (0.010)	0.007 (0.010)	0.011 (0.012)	0.016 (0.013)
<i>Employment</i>						
T*post2011	0.030*** (0.011)	-0.001 (0.011)	-0.003 (0.009)	0.007 (0.010)	0.007 (0.009)	0.004 (0.016)
T*post2011*high lambda	-0.006 (0.014)	0.012 (0.013)	0.011 (0.011)	-0.016 (0.012)	-0.017 (0.012)	-0.015 (0.014)
N	2453	3332	3090	3576	3856	3193

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .



# Effects on income (no self-employed)

	Women				Men	
	55-59	50-54	45-49	55-64	50-54	45-49
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Labour income</i>						
T*post2011	461.679*** (172.064)	-117.927 (165.291)	156.931 (119.276)	37.509 (289.677)	-133.892 (277.224)	-275.212 (447.801)
<i>Total income</i>						
T*post2011	578.846*** (219.123)	-17.615 (190.717)	138.773 (143.599)	-103.342 (335.452)	-275.824 (406.235)	136.327 (597.895)
<i>Savings</i>						
T*post2011	752.686* (389.635)	273.928 (220.401)	139.661 (231.913)	382.444 (272.291)	-317.404 (336.174)	280.112 (322.087)
N	1831	2730	2548	2708	2967	2470

**Notes:** Additional controls: year and cell  $q$  fixed effects, region and sector fixed effects, time fixed effects, marital status. Only individuals with at least 10 and less than 40 accrued years of contribution. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

# More results on employment in INPS

	Employed	wage growth	full time	part time	temp contract	perm contract	change firm	av wage firm
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Women								
T*post2011	0.005*** (0.001)	-0.000 (0.001)	0.002** (0.001)	0.003*** (0.001)	-0.001 (0.001)	0.004*** (0.001)	0.001 (0.001)	0.005*** (0.001)
N	485180	174904	485093	485180	190719	485180	173849	190695
Men								
T*post2011	0.000 (0.001)	-0.001 (0.001)	-0.002 (0.001)	0.002*** (0.000)	-0.001 (0.001)	-0.002 (0.001)	0.004*** (0.001)	-0.001 (0.002)
N	822984	414407	822977	822984	446020	822984	396664	445700

Notes: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .



# Descriptive statistics at the family level

	Family level analysis							
	All wife 45-59 [1]	Treated Wives Sample not elig [2]	Control $T_q^w < 7$ [3]	Treated $T_q^w \geq 7$ [4]	All husb 45-64 [5]	Treated Husbands Sample not elig [6]	Control $T_q^h < 4$ [7]	Treated $T_q^h \geq 4$ [8]
Age w	51.654 (4.264)	51.761 (4.026)	51.051 (3.781)	52.349 (4.127)	50.240 (6.752)	50.499 (3.663)	50.240 (3.669)	50.846 (3.628)
Age h	55.089 (5.702)	55.286 (5.149)	54.524 (4.926)	55.917 (5.245)	54.052 (5.682)	52.984 (3.912)	52.545 (4.145)	53.569 (3.494)
Y. contrib w	15.128 (13.447)	24.715 (7.663)	30.813 (4.204)	19.668 (6.036)	14.581 (13.519)	24.434 (7.492)	24.584 (7.522)	24.235 (7.449)
Y. contrib h	30.286 (10.517)	33.839 (6.415)	33.934 (5.826)	33.759 (6.864)	29.026 (10.698)	30.734 (4.486)	31.651 (4.708)	29.510 (3.846)
High edu w	0.486	0.603	0.672	0.545	0.501	0.664	0.615	0.729
High edu h	0.483	0.558	0.599	0.523	0.493	0.627	0.548	0.732
If children	0.705	0.690	0.665	0.712	0.678	0.654	0.654	0.654
Active w	0.523	0.841	0.971	0.733	0.508	0.859	0.838	0.886
Active h	0.768	0.777	0.820	0.741	0.794	0.987	0.984	0.991
Unempl w	0.035	0.039	0.019	0.055	0.037	0.037	0.037	0.036
Unempl h	0.058	0.025	0.020	0.030	0.060	0.030	0.034	0.024
Log(wage net) w	9.510 (0.592)	9.563 (0.555)	9.662 (0.447)	9.435 (0.646)	9.490 (0.608)	9.583 (0.546)	9.542 (0.554)	9.633 (0.531)
Log(wage net) h	9.841 (0.470)	9.927 (0.420)	9.969 (0.379)	9.886 (0.453)	9.829 (0.461)	9.933 (0.420)	9.920 (0.415)	9.948 (0.425)
Observations	11843	5510	2566	2944	15204	3825	2166	1659

# Couples, full specification interacting for husbands' eligibility

Couples where:	wives not eligible to retire 2010			husb. not eligible to retire 2010		
	Active husb. (1)	Unemp. husb. (2)	Employed husb. (3)	Active wife (4)	Unemp. wife (5)	Employed wife (6)
T wife*post2011	0.015* (0.008)	0.007 (0.005)	0.007 (0.008)	0.006 (0.006)	0.010** (0.004)	-0.004 (0.006)
T wife*post*elig husb	0.012 (0.021)	-0.005 (0.012)	0.016 (0.021)			
T husb*post2011	-0.011 (0.010)	-0.002 (0.010)	-0.009 (0.014)	-0.011 (0.010)	-0.008 (0.006)	-0.003 (0.011)
T husb*post*elig husb	0.160*** (0.047)	0.054** (0.024)	0.107** (0.048)			
N	5326	5326	5326	3819	3819	3819

**Notes:** Most of the husbands' response to their wives' perspective effect comes from husbands already eligible to retire, who decide to postpone retirement. T wife\*post 2011 is the estimated difference-in-differences coefficient of the wife's longer working horizon (perspective effect); T husband\*post 2011 is the estimated difference-in-differences coefficient of the husband's longer working horizon (both direct and perspective effect). Additional controls: year and cell  $q_s$  and  $q_s$  fixed effects (separately for each dimension; each cell is defined by age, gender number of years of accrued contribution and sector of employment), region and sector fixed effects, age difference across partners (also squared) and difference in distance to retirement across partners (also squared), partner's change in distance to retirement. The sample of columns 1-3 (4-6) consists of all husbands (wives) older than 45, belonging to couples where the wives (husbands) are not eligible for a public pension either before and after the reform and have accrued a number of years of contribution of at least 10, and smaller than 40. Elig is a dummy equal to one if the husband is eligible at time  $t$  to a public pension. Columns 4 to 6 do not distinguish whether the wife is eligible for a public pension or not because there are too few wives eligible for a public pension whose husband is in the sample (is not eligible for a public pension).

Robust standard errors clustered at the cell  $q_s$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

# Couples, results by education level


	Shock to wife MRA on wife (1)	Shock to wife MRA on husband (2)	Shock to husband MRA on husband (3)	Shock to husband MRA on wife (4)
<i>Participation</i>				
T wife*post2011	0.047*** (0.007)	0.062*** (0.022)		
T husb*post2011			-0.005 (0.005)	-0.070 (0.060)
T wife*post 2011*own high edu	-0.023*** (0.008)	-0.066** (0.030)		
T husb*post 2011*own high edu			0.002 (0.006)	0.103 (0.069)
<i>Unemployment</i>				
T wife*post2011	0.023*** (0.006)	0.002 (0.016)		
T husb*post2011			-0.010 (0.013)	-0.010 (0.035)
T wife*post 2011*own high edu	-0.011* (0.006)	-0.010 (0.023)		
T husb*post 2011*own high edu			0.009 (0.014)	-0.004 (0.039)
<i>Employment</i>				
T wife*post2011	0.024*** (0.008)	0.060** (0.029)		
T husb*post2011			0.005 (0.014)	-0.060 (0.064)
T wife*post 2011*own high edu	-0.012 (0.009)	-0.056 (0.036)		
T husb*post 2011*own high edu			-0.006 (0.015)	0.107 (0.074)
N	5142	5142	3646	3646

**Notes:** Additional controls: year and cell  $q_s$ , and  $q_s$  fixed effects, region and sector fixed effects, age difference across partners (also squared) and difference in distance to retirement across partners, partner's change in distance to retirement. The sample in columns 1 and 2 (3 and 4) consists of couples where the wives (husbands) are not eligible for a public pension either before and after the reform and have accrued more than 9 (19) and less than 40 years of contribution. High edu is a dummy equal to 1 at

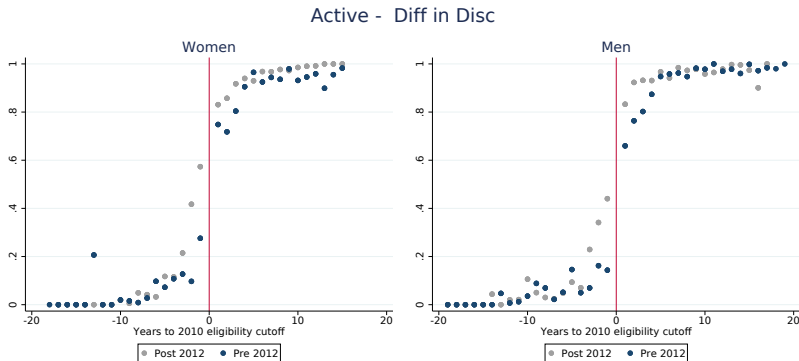
# The direct eligibility effect - estimating equation

$$Y_{it} = post2011_t [E_i(\beta + f_r(d_{2010i})) + (1 - E_i)f_l(d_{2010i})] + E_i(\delta + f_r^P(d_{2010i})) + (1 - E_i)f_l^P(d_{2010i}) + \psi_t + v_{it}$$

- $Y_{it}$  represents labour supply of individual  $i$  in year  $t$ ;
- $post2011_t$  is an indicator for the post reform period;
- $E_i$  indicates whether individual  $i$  was eligible to retire under the pre-reform rules;
- $d_{2010i}$  is distance to retirement under the pre-reform rules (our running variable);
- $f_r$ ,  $f_l$ ,  $f_r^P$  and  $f_l^P$  are some polynomials of  $d_{2010i}$ ;
- $\psi_t$  represent year fixed effects and  $v_{it}$  is an error term, which we cluster at the individual level

Only for women in the control group (not affected by perspective effect off of the eligibility threshold) 

# Diff-in-disc: graphical representation for individuals



**Source:** SHIW, from 2010 to 2016.

**Note:** The Figure shows the evolution of the probability of being active in the labour market as a function of the distance to the minimum retirement age according to the 2010 pension rules (that is the difference between the minimum retirement age under 2010 rules and the age at the interview). The blue (grey) dots are observations in years pre-2012 (post-2012). The Figure shows that after 2012 individuals eligible to retire under the previous pension rules have a higher probability of being active in the labour market and that the discontinuity around zero (which means that the individual would have reached the pension eligibility according to the 2010 rules) does not longer hold.

# At the individual level: effects both on women and on men

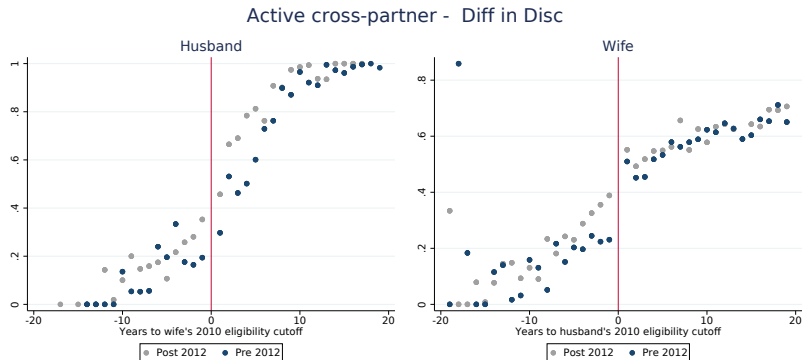
	linear (1)	Women quadratic (2)	cubic (3)	linear (4)	Men quadratic (5)	cubic (6)
<b>Individual effect</b>						
<i>Participation</i>						
$E_{2010}^{*post2011}$	0.213*** (0.037)	0.273*** (0.052)	0.250*** (0.068)	0.105*** (0.027)	0.101*** (0.039)	0.081 (0.051)
<i>Unemployment</i>						
$E_{2010}^{*post2011}$	0.027** (0.013)	0.007 (0.021)	-0.021 (0.032)	0.002 (0.012)	0.006 (0.020)	-0.047 (0.030)
<i>Employment</i>						
$E_{2010}^{*post2011}$	0.185*** (0.038)	0.266*** (0.055)	0.270*** (0.075)	0.103*** (0.028)	0.095** (0.042)	0.128** (0.057)
N	9902	9902	9902	18784	18784	18784

**Notes:** Additional controls: year fixed effects. The sample consists of individuals whose distance to retirement according to the pre-reform rules ( $d_{2010}$ ) was between 20 and -20. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

Women's response is about four times larger at the eligibility threshold 



# Diff-in-disc: graphical representation for couples



**Source:** SHIW, from 2010 to 2016.

**Note:** The Figure displays the probability of being active as a function of own partner's distance to retirement according to the 2010 pension rules. The blue (grey) dots are observations in years pre-2012 (post-2012). It shows that after 2012 partners of those individuals would have been eligible to retire according to the 2010 rules do not have a higher probability of being active.

# The direct eligibility effect: no effects across partners

	on husbands (wives' shock)			on wives (husbands' shock)		
	linear	quadratic	cubic	linear	quadratic	cubic
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Participation</i>						
$E_{2010}^{s'} \text{ *post2011}$	-0.022 (0.036)	-0.044 (0.054)	-0.031 (0.075)	0.065** (0.028)	0.012 (0.042)	-0.031 (0.059)
<i>Unemployment</i>						
$E_{2010}^{s'} \text{ *post2011}$	0.013 (0.016)	0.001 (0.024)	0.030 (0.032)	-0.011 (0.010)	-0.004 (0.015)	-0.011 (0.021)
<i>Employment</i>						
$E_{2010}^{s'} \text{ *post2011}$	-0.034 (0.037)	-0.045 (0.055)	-0.062 (0.077)	0.076*** (0.029)	0.017 (0.042)	-0.020 (0.057)
N	5924	5924	5924	14375	14375	14375

**Notes:** Additional controls: year fixed effects. The sample consists of individuals whose distance to retirement according to the pre-reform rules ( $d_{j/2010}$ ) was between 20 and -20. Robust standard errors clustered at the cell  $q$  level in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .